- Grades and comments on your Du Bois figures are up on Blackboard
- Grades for the first referee report will be up before Spring Break, email me if you did not get a confirmation
- Remember that the second referee report is due March 15th on Miller (2008) "Women's Suffrage, Political Responsiveness, and Child Survival in American History"
- Also think about making progress on your data projects
- No class on the Friday before Spring Break

- For the Clark and Cummins paper, check out the Stata and Excel examples on Blackboard for more on the issue of underlying social capital
- Now back to Clark's approach to the US
- Clark's sources for elite groups will be:
 - Descendants of Ashkenazi and Sephardic Jews
 - Descendants of wealthy individuals in 1923-24 with rare surnames
 - Descendants of individuals with rare surnames graduating from Ivy League schools in and before 1850
- Note how much more limited the time range needs to be

• Clark's sources for underclass groups will be:

- Native Americans
- Black Americans whose ancestors came to the United States before the Civil War
- Descendants of the French settlers who came to the colonies between 1604 and 1759
- Think about how these groups differ from those used by Clark for Britain

- Measuring outcomes requires a different approach as well for the US
- Probate records are not as easily accessible (you would have to do a lot of work requesting one record at a time from many different locations)
- Instead, Clark is going to take an approach similar to looking at Cambridge and Oxford graduates
- He'll take advantage of the public directories of doctors and lawyers

Abbeville

ALABAMA

Attalla

PHYSICIANS OF ALABAMA

Capitals signally memorranty in Courtiz and State Societies. It membership in American Medical Association ; H, Amongashi F, celettici PM, physionemics ; Information recarding graduation not obtained; (V, of P.) in pratice before passes of present have or interaction of postof a certain number of years of practice; I. S., United States Indian Strivic; M. H.C., United States Medical Boerce Courts; increasing and the state of the states of the state of the states of the state of the states of the state of the states of the states of the states of the state of the states of the sta

ALABAMA.

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RADFORD, GEO, CLEMENTS-* (1)587); R.F.D. No. 2 REAGAN, ONSLOW-* (1 1882). STREET, THOMAS HEZEKIAH (b 1876) -Pa.2,'00; (1 1900); office, Rainey Bldg. Alexandria, 63, Calhoun. Meharg, R. L .- Ala.2.'06: (1 1906). Aliceville, 350, Pickens, Duncan, Wallace W.-Ala.4,'99; (1 1899), Moody, Joseph (b 1846)-Ky.1, 71; (1 1878). Murphy, Chas. M .- Ala.4.* (1 1). Snoddy, Ephraim A .- Ala.2,'97; (1 1897). Allenton, 210, Wilcox. Lee, John F.-Tenn.1,'80; (11883). Alpine, 121, Talladega, HARGROVE, ROBT. HARRIS - Tenn.5.'87: (1 1887). REEVES. THOMAS EDWIN-Tenn.11,'06; (1 190€). SMITH, FRANK C. (b 1879) - Ala.4,'03; (1 1903). Aiston, 65, Barbour (R.F.D., Blue Springs). BUSH, DAVID-Ala, "07: (1 1907). Alton, -, Jefferson. HUEY, J. F.-Md.3,'96; (1 †). Altoona, 1,000, Etowah, ELLISON, JOHN HENRY-Tenn.5,'88;(1 1899). America, 100, Walker. Jones, Giles Wilson (b 1874)-Tenn.9,'00;

(1 t).

Warren, Wm, James-Ga.5,'89; (11889); not in practice. White, Wm. Y .- Tenn.5,'87; (1 1887). WHITESIDE, JOHN MCINTYRE (b 1862) -Tenn.5,'94; (1 1894); Queenland Ave. and 7th St.; office, Scarborough Drug Co. Wilke, Jesse Lane-Ga.1, 71; (11551); not in practice. WILLIAMS. BENJ. DUDLEY (b 1854)-La.1, 'S1; (1 1887). WILLIAMS, MARK JOHNSON⊕-Ala.4.'02; (1 1902). WINN, LOCKE MINOR-La.1,'00; (1 1900). Ansley, 95, Pike. BROACH, FRANCIS MARION (b 1855)-Ga.10. '90: (1 1890) Dennis, S. H. (b 1853) - Ky.4, 58; (1 1878); R.F.D. Arab, 300, Marshall. Hinds, Montgomery L .-- Tenn.5, '91; (1 1892), Hinds, Wm.-Ala.2,* (1 1). Ariton, 250, Dale. NORRIS, ROY HART-Ala.2, 97; (1 1897) PATTON. JOHN HAMPTON-Md.9, '02; (1 1902). Weed, Saml, Lafavette-Ala.2.'86; (1 1887). Arkadelphia, 203, Cullman. LEE, GENERAL ROBT.-Ala.4.'06: (1 1996). PARKER, D. J .- Tenn. 8, '01; (1 1901). Arley, 75, Winston,

DENNIS, DAVID R. (b 1840)-* (1 1902).

February 25, 2019 5 / 45

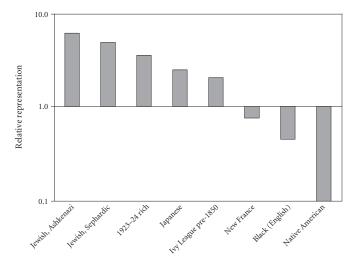


FIGURE 3.1. Relative representation of surname types among physicians.

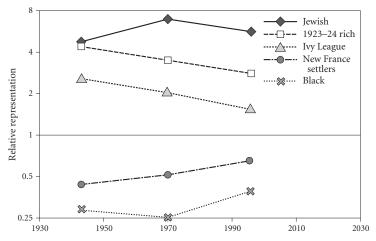


FIGURE 3.4. Relative representation of surname types among physicians, by generation.

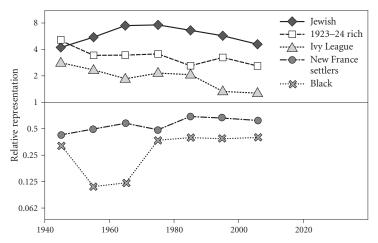


FIGURE 3.8. Relative representation of surname types among physicians, by decade.

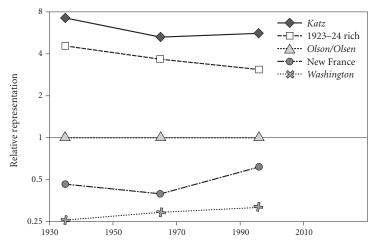


FIGURE 3.10. Relative representation of surname type among attorneys, by generation.

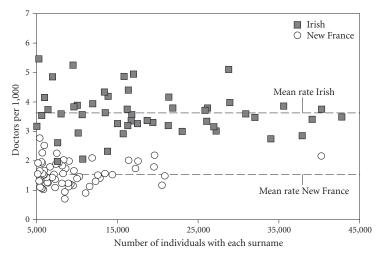


FIGURE 3.11. Physicians per thousand surname holders, most common Irish and New France surnames.

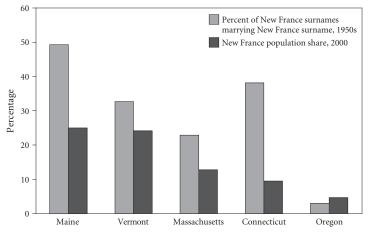


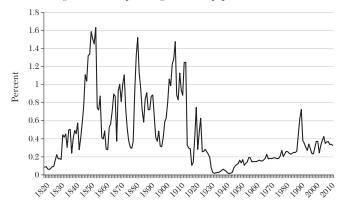
FIGURE 3.12. Marital endogamy among New France descendants, 1950s.

Immigration and Mobility

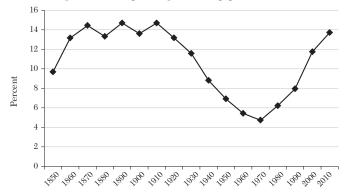


Immigration and Mobility





Panel A. Forign-born flow as percentage of the US population (1820-2010)



Panel B. Forign-born stock as percentage of the US population (1850-2010)

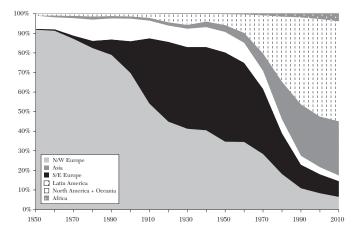


Figure 2. Sending Regions within the Foreign-Born Population, 1850-2010

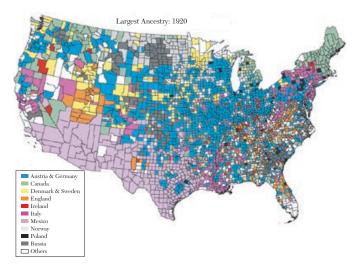
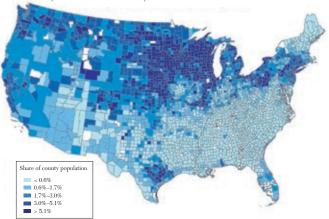


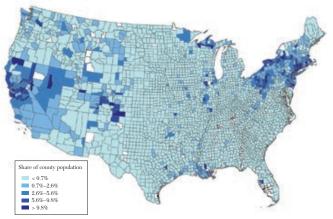
Figure 3. Largest Country-of-Origin Group among Foreign Born by County, 1920

American Mobility, Spring 2019

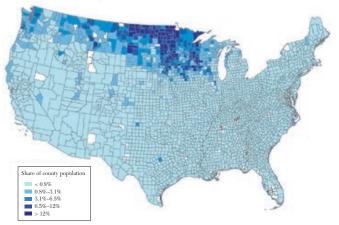


Panel A. Ancestry Share: Austria and Germany, 1920

Panel B. Ancestry Share: Italy, 1920



Panel C. Ancestry Share: Norway, 1920



Immigration in the Colonial Period

Cherryland, Cal new of Hath put himfelf, and by their Prefents, dech voluntarily, and of his own free Will and Acceed, per bindfill Apprendic to der voltrandly, and of hig over the Will and Acceed, per bindfill Apprendic to der softward and the bases ha Ar., Trade, and Medray, and she the Massee of an Apprendic to force the constraint and the softward for from the Dyr of the Data Hered, for, and altering, and much the full End and Terms of converse for thereast, for and altering, and much the full End and Terms of converse for the softward of the in Spendle - next enforme. During all which Term, the faid Apprentice his faid Mafter faithfully thall ferve, his Socrets keep, his lawful Commands everywhere readily obey. He fhall do no Damage to his faid Mather, nor fee it to be done by others without loxing or giving Notice thereof to his faid Mafter. He thall not watte his faid Mafter's Goods, nor lead them unlawfully to any, He fhall not commit Furnication, nor contraft Matrimouv within the fald Term, At Cards, Dice, or any other unlawful Game, he thull not play, whereby his faid Mafter may have Damage. With his own Goods, nor the Goods of others, without Licence from his faid Mafter, he thall neither buy nor fell. He thall not ablent himfelf Day nor Night from his faid. Mafter's Service, without his Leave : Nor haunt Alc-boofes, Taveras, or Play-houfes ; but in all Things behave himfelf as a faithful Appenntice ought to do, during the faid Term. And the faid Mather fhall ufe the utmost of his Endeavour to reach or cause to be mught or inftructed the faid Apprentice in the Trade or Muffery of to be under ar inheled the field Argenetic in the Trade or Mythery of The Control of the second second second second second second second Marine, Dirich, & Control of the Control of the Second sec AND for the true Performance of all and fingular the Covenants and Agreed ments aborefaid, the faid Parties bind themfelves each unto the other family by mens aberlaid, the taid puries total dentifieses each usin no contex tensity op-thed Preters. In WUTNESS between the fail the Third Jarre interchange-ably for them Hands and Sanka horseans. Data the de-Day of *exceedings* is the *discretionally*. Year of the Reign of our Scoretting Local Linguist for exceeding the State horizon, for desayer Down One Thouland Score Handsen and Scoretty. Scoled and delivered in the Profesce of as and a rorant los Civistop lan chargeson Richand Berguson

This indentured...between [Alexander Beard]...of the one part, and [John Dickey]...of the other part, witnesseth, that the said [Alexander Beard] doth hereby covenant, promise and grant, to ...[John Dickey]...and his assigns, from the day of the date hereof until the first and next arrival at [Philadelphia] in America...and during the term of [three] years to serve in such service and employment as the said [John Dickey] or [his] assigns shall there employ [him]...In consideration whereof the said [John Dickey] doth grant...to pay for [his] passage, and to find allow [him] meat, drink, apparel and lodging, with other necessaries, during the said term; and at the end of the said term to pay unto him the usual allowance, according to the custom of the country in the like kind ...

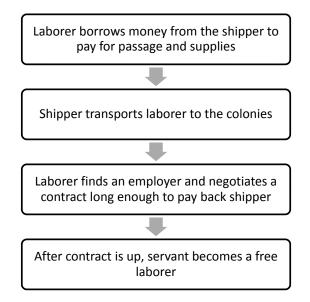
Credit Constraints and Indentured Servitude

- The cost of passage to America was $\pounds 5$ to $\pounds 10$, an amount greater than average annual income at the time
- To put that in perspective, think about college tuition:
 - Average tuition and fees at private four-year colleges is \$32,410 (according to the College Board)
 - Median income for a 20 to 24 year old is \$26,728 (according to the BLS)
 - If there were no student loans, how would people pay for college?

How Indentured Servitude Works - Standard



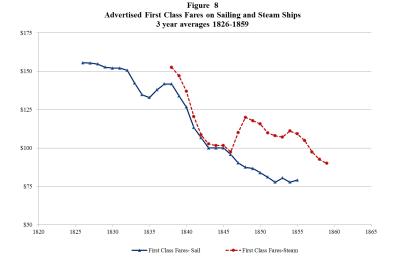
How Indentured Servitude Works - Redemptioners



Migration and Human Capital During the Colonial Period

Contract Length and Servant	
Characteristics	
	Months More or
Characteristic	Less Service
15 years old	26
17 years old	9
19 years old	2
Female	-2
Literate	-1
Farmer	-4
Metalworker	-4
Textile worker	-4

.









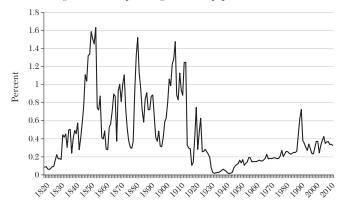
Chinese Exclusion Act - 1882



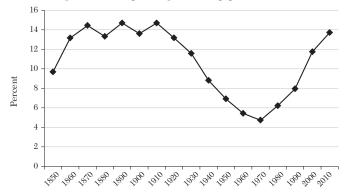
Chinese Exclusion Act - 1882



- Grades will be posted tomorrow for your referee reports, I will email each of you comments on your report
- Remember that the second referee report is due March 15th on Miller (2008) "Women's Suffrage, Political Responsiveness, and Child Survival in American History"
- This is a different style of paper, be certain to think about whether they are proving their hypothesis or whether there are alternative explanations
- Also think about making progress on your data projects, feel free to email me if you run into any problems
- No class on Friday, have a great Spring Break



Panel A. Forign-born flow as percentage of the US population (1820-2010)



Panel B. Forign-born stock as percentage of the US population (1850-2010)



Chinese Exclusion Act - 1882



Chinese Exclusion Act - 1882

Ind	ustrial Distribu	tion of Chinese E	mployment by Re	egion, 1870-193	30
	Total	Northeast	Midwest	South	West
		<u>18</u>	70		
Restaurants	0.2				0.2
Laundries	11.0				11.0
Food stores	1.3				1.3
All else	87.5				87.5
		18	80		
Restaurants	0.4	0.0	0.0	0.0	0.4
Laundries	13.9	100.0	100.0	0.0	11.9
Food stores	1.6	0.0	0.0	0.0	1.6
All else	84.1	0.0	0.0	100.0	86.1
		190	00		
Restaurants	0.5	0.0	0.0	0.0	0.6
Laundries	34.1	81.6	100.0	85.7	16.9
Food stores	3.4	0.0	0.0	7.1	4.2
All else	62.0	18.4	0.0	7.3	78.3
		19	10		
Restaurants	7.4	9.2	11.1	8.5	5.6
Laundries	20.9	60.2	66.7	31.9	7.4
Food stores	6.2	0.0	0.0	31.9	12.7
All else	65.5	30.6	22.2	27.7	74.3
		192			
Restaurants	17.3	32.8	47.8	22.2	12.2
Laundries	22.1	55.5	39.1	44.4	11.5
Food stores	7.4	5.5	0.0	7.4	7.0
All else	53.2	6.2	13.1	26.0	69.3
		19			
Restaurants	27.7	42.0	32.4	34.4	15.6
Laundries	24.7	42.0	50.0	21.9	6.6
Food stores	8.8	0.6	0.0	28.1	13.2
All else	38.8	15.4	17.6	15.6	64.6

J. Parman (College of William & Mary)

American Mobility, Spring 2019

Distribution of the Chinese-American Population, 1870-1960						
	Percentage of counties	Median number of Chinese				
	with one or more Chinese	residents in a county with				
Year	residents	Chinese residents				
1870	10.9	19				
1880	18.8	1				
1890	37.8	4				
1900	45.7	4				
1910	40.8	5				
1920	44.9	4				
1960	42	7				

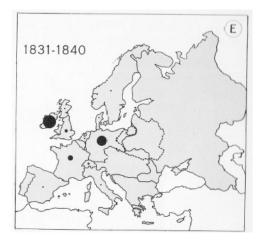
Dissimilarity in Residence by State						
	Chinese vs	Black vs Non-				
	Non-Chinese	Black				
1870	97.3	66.6				
1880	93.6	66.6				
1890	85.4					
1900	70	66.2				
1910	65.7	67.7				
1920	55.6	49.1				
1930	57.4	54.5				

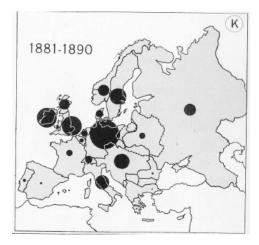
Sec. 3. That the following classes of aliens shall be excluded from admission into the United States: All idiots, imbeciles, feeble-minded persons, epileptics, insane persons...persons of constitutional psychopathic inferiority; persons with chronic alcoholism; paupers; professional beggars; vagrants; persons afflicted with tuberculosis... ...persons who have been convicted of or admit having committed a felony or other crime or misdemeanor involving moral turpitude; polygamists; anarchists...[persons] who advocate or teach unlawful destruction of property; ...persons coming to the United States for the purpose of prostitution or for any other immoral purpose... ...[The provision] shall not apply to the persons of the following status or occupations: Government officers, ministers or religious teachers, missionaries, lawyers, physicians, chemists, civil engineers, teachers, students, authors, artists, merchants, and travelers for curiosity or pleasure... All aliens over sixteen years of age, physically capable of reading, who can not read the English language, or some other language or dialect, including Hebrew or Yiddish...That for the purpose of ascertaining whether aliens can read the immigrant inspectors shall be furnished with slips of uniform size...each containing not less than thirty nor more than forty words in ordinary use, printed in plainly legible type of some one of the various languages or dialects of immigrants.

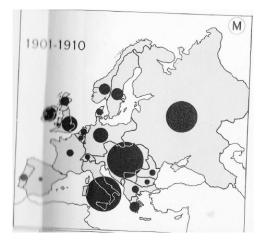


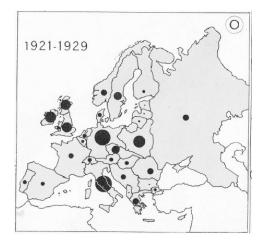
1917 - Literacy Test











Immigration and Nationality Act - 1965



- You should have gotten an email with my feedback on your first referee report, let me know if you didn't
- Remember that the second referee report is due March 15th on Miller (2008) "Women's Suffrage, Political Responsiveness, and Child Survival in American History"
- This is a different style of paper, be certain to think about whether they are proving their hypothesis or whether there are alternative explanations
- It's also time to get moving on your data projects, feel free to email me if you run into any problems
- On Wednesday we'll talk about your final projects



Economics 449: American Economic Mobility Over Two Centuries

This site displays the original work of Econ 449 students exploring the history of mobility and inequality in the Williamsburg through a variety

The beta version of our class website is up (with your data still to come). Suggestions are welcome. Let me know if you want your name included with your figure.

half a century ago. Deed histories

shaped the evolution of those

neighborhoods.

uncover the restrictive covenants that

occupational and geographical mobility

over a time period that witnessed

tremendous shocks to the national

economy as well as major changes specific to the Williamsburg community.

resulting figures provide insights into the

ways racial inequalities varied across the

South and evolved over the twentieth

century.

- The outcomes of immigrants will depend crucially on which individuals decide to migrate
- In particular, it will depend on whether there is *positive* or *negative* selection into migration
- Economists have developed a theory of selection into migration that has its roots in Roy's 1951 article "Some Thoughts on the Distribution of Earnings"
- The keywords for the article: hunting, rabbits, fishers, occupations, productivity, trout, logarithms, communities, industrial productivity, relative prices

Selection Into Migration

- While Roy may have started with rabbits and trout, Borjas formalized the Roy model for immigration with his 1987 article "Self-Selection and the Earnings of Immigrants"
- Let's discuss the basics of the model (see Autor's notes for more details)
- There are two countries, 0 (the source country) and 1 (the host country)
- Log earnings in the source country are given by:

$$w_0 = \mu_0 + \varepsilon_0$$

• Log earnings in the host country are given by:

$$w_1 = \mu_1 + \varepsilon_1$$

• The log earnings equations:

$$w_0 = \mu_0 + \varepsilon_0$$

$$w_1 = \mu_1 + \varepsilon_1$$

- μ_i designates the mean log earnings in country *i*
- ε_i can be thought of as the returns to a worker's skill relative to the mean worker in country i
- Assume that ε_i is distributed normally with mean zero and variance σ_i^2

- It is costly to migrate
- Let's assume for the moment that $\mu_0 \approx \mu_1$ (but do spend some time thinking about alternative cases)
- For migrants, the gains from migration in terms of switching from ε₀ to ε₁ have to outweigh these costs
- So does this mean that migrants will necessarily be very high ability?
- Maybe, maybe not
- The nature of selection will depend crucially on σ_0 , σ_1 and ρ (the correlation between σ_0 and σ_1)

Case 1: Immigrants are positively selected from the source country and are above the mean of the host country $(E(\varepsilon_0|migrant) > 0 \text{ and } E(\varepsilon_1|migrant) > 0)$

- This will be the case if $\frac{\sigma_1}{\sigma_0}>1$ and $\rho>\frac{\sigma_0}{\sigma_1}$
- The first condition says that the return to skill is greater in the host country than the source country
- The second condition essentially says that being high skilled in the source country will translate into being high skilled in the host country

Case 2: Immigrants are negatively selected from the source country and are below the mean of the host country $(E(\varepsilon_0|migrant) < 0 \text{ and } E(\varepsilon_1|migrant) < 0)$

- This will be the case if $\frac{\sigma_1}{\sigma_0} < 1$ and $\rho > \frac{\sigma_1}{\sigma_0}$
- The first condition says that the return to skill is greater in the source country, so low skilled workers would prefer the compressed wage distribution of the host country
- The second condition essentially says that being low skilled in the source country will translate into being low skilled in the host country

Case 3: Immigrants are selected from the lower tail of the source country and end up above the mean of the host country $(E(\varepsilon_0|migrant) < 0 \text{ and } E(\varepsilon_1|migrant) > 0)$

• This can occur if
$$ho < min\left(rac{\sigma_1}{\sigma_0}, rac{\sigma_0}{\sigma_1}
ight)$$

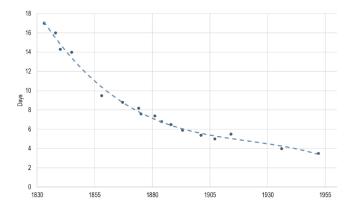
- This is effectively saying that there is some reason that a migrant receives relatively low pay in the source country but ends up highly paid in the host country
- When could this happen?

- Let's look at some empirical evidence of selection into migration during the age of mass migration
- It is an interesting time period to study because of the nearly open border
- The central question is whether immigrants were positively or negatively selected
- We'll look at empirical evidence of both cases
- After the question of selection is addressed, there is a second question: what sort of mobility did immigrants experience after arriving in the United States?

- First, how do we look at who decided to emigrate?
- One classic approach is to turn to passenger lists
- In 1819, the US starts requiring ship captains to file ship records upon arrival in the United States
- Several economic history studies have used these passenger lists to get a sense of the age and occupational distributions of emigrants

..[E]very ship or vessel bound on a voyage from the United States to any port on the continent of Europe, at the time of leaving the last port whence such ship or vessel shall sail, shall have on board, well secured under deck, at least sixty gallons of water, one hundred pounds of salted provisions, one gallon of vinegar, and one hundreds pounds of wholesome ship bread, for each and every passenger on board such ship...

Passenger Act of 1819



Liner transatlantic crossing times, 1833-1952 (P.J. Hugill (1993))

Passenger Act of 1819

...[T]he captain or master of any ship or vessel arriving in the United States, or any of the territories thereof, from any foreign place whatever...shall also deliver and report, to the collector of the district in which such ship or vessel shall arrive, a list or manifest of all the passengers taken on board of the said ship or vessel at any foreign port or place; in which list or manifest it shall be the duty of the said master to designate, particularly, the age, sex, and occupation, of the said passengers, respectively, the country to which they severally belong, and that of which it is their intention to become inhabitants: and shall further set forth whether any, and what number, have died on the voyage;

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- So what can we do with passenger lists?
- The most basic thing is to look at the gender and age distributions of individuals arriving in the US
- Some passenger lists include occupation, although there is a (very interesting) question of how that will correspond to occupation in the US
- One thing we would really like if we're focusing on selection is a measure of human capital

Passenger Lists and Human Capital

- The manifests do not contain a direct measurement of education level beyond (sometimes) literacy
- If we don't know education level or literacy, what else can we use?
- It turns out that a decent source of information on human capital is actually age
- We can infer something about the basic human capital of immigrants by looking at the distribution of ages, specifically the distribution of the last digit of age
- Let's head over to Stata to see how this works

Passenger Lists and Human Capital

- Mokyr and O'Grada use this concept of age heaping to look at Irish immigrants
- They use a slightly more complex measure than what we just saw
- In particular, they are measuring

$$\gamma = \sum_{i=15}^{34} \left(\frac{n_i}{\sum n_i} - \frac{\hat{n}_i}{\sum \hat{n}_i} \right)^2$$

• This is basically saying how much deviation do we see from a smooth age distribution

Passenger Lists and Human Capital

Population	Males	Females	Total
Ireland, 1841	54.18	72.36	66.09
Ulster, 1841	47.46	61.54	54.49
Leinster, 1841	43.98	66.40	54.80
Munster, 1841	57.72	88.21	72.32
Connaught, 1841	80.14	117.70	98.98
Ireland, 1851	63.23	89.24	88.42
Ulster, 1851	53.52	57.89	62.33
Leinster, 1851	50.00	61.80	74.88
Munster, 1851	72.13	115.50	92.27
Connaught, 1851	97.00	117.53	140.44
U.S. 1880			10.23
U.S. 1970			1.16
Mexico, 1960			22.72

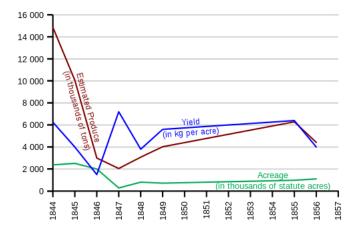
Values of γ based on census data

Values of γ based on passenger lists

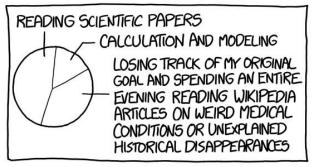
Sample	Males	Females	Total
1803-06, Ireland	89.8	130.8	96.3
1819-20, Ireland	146	108.5	106.5
1830-39, Derry	262.9	319	268.5
1820-48, New York	113.2	158.7	122.9
1822-39, Boston	125.9	117	111.9
1820-50, Sweden	22.1	27.9	20.5

Values of $\boldsymbol{\gamma}$ based on New York passenger lists only

Period	Males	Females	Total
1820-24	103.6	209.8	107.4
1825-29	113.1	170.2	114.2
1830-34	96	109.5	96.7
1835-39	100.5	143.9	104.9
1840-44	92.2	141.5	103.8
1845-46	158.7	175.1	150
1847-48	217	296.3	239.9
Total	113.2	158.7	112.9



MY RESEARCH METHODS:



The Basics

- You need to do an original empirical study related to local mobility or inequality in a historical context
- You may want to focus on changes in mobility or inequality between groups over time
- You may also want to consider the impact of a historical shock on mobility or inequality
- Either way, you need to do original research that includes:
 - Developing a research question with a testable hypothesis
 - Collecting micro-level data (you can use the class datasets, you can also decide to use other data sources)
 - Presenting summary statistics and stylized facts using those data
 - Presenting novel regression analysis to test your hypothesis
- You will produce both a research paper and a policy

memo

J. Parman (College of William & Mary)



What is historical?



American Mobility, Spring 2019

- The research paper should be approximately 15 pages double-spaced inclusive of figures and references
- Write it in the style of an economics journal article
- Typical components:
 - Introduction
 - Literature review
 - Description of data and methods
 - Presentation of empirical analysis
 - Conclusion
- Key thing is that this requires original data analysis

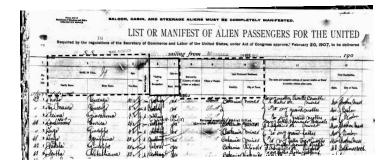
- Two pages double-spaced with at least one figure or table (this is a strict limit)
- It should contain:
 - A succinct statement of the policy issue
 - A concise summary of your methodology and findings
 - Discussion of the importance of your findings to policy
 - Recommendation for policy change
- Must be written for non-economists (and non-locals)

- Both the paper and policy memos are due by 5pm on April 29th
- Submit them by email to me (jmparman@wm.edu)
- They should be submitted as separate pdf's (both can be sent in the same email)
- Pay attention to formatting, it does matter for the grading
- I'm happy to look over drafts and give feedback

Every failure is a step to success. Every detection of what is false directs us towards what is true: every trial exhausts some tempting form of error. Not only so; but scarcely any attempt is entirely a failure; scarcely any theory, the result of steady thought, is altogether false; no tempting form of Error is without some latent charm derived from Truth. – W. Whewall, Lectures on the History of Moral Philosophy in England, 1852

Some Quick Tips

- Get started early, even if you plan to use the class data (you know what variables will be in there)
- If using Stata, take advantage of the tutorials on our course website and the online help files and forums
- Use my office hours, especially if you are getting stuck on something technical
- Also consider meeting with the research librarians
- Check out the resources in the final project guidelines on Blackboard (and that we covered when talking about the Du Bois project)
- Google Scholar and Google Books are your friends



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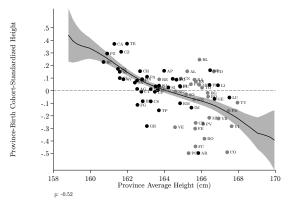


Figure 3: Province-cohort z-score and average province height.

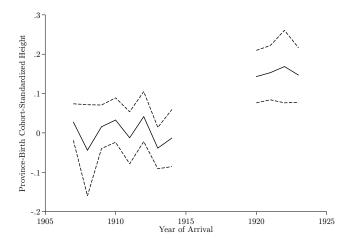


Figure 4: Local selection by year of arrival.

- So height gives us another way to get at human capital
- Height does suffer from some limitations (genetic differences across different population groups, inability to capture variation after full height is achieved, ...)
- One thing it does show us in the context of Spitzer and Zimran's article is that the nature of selection can vary across localities within a sending country

- Let's switch to something familiar, Long and Ferrie's work on the US versus Britain
- We've already looked at how Long and Ferrie could link individuals within the US and within Britain
- This strategy also works for linking individuals between the US and Britain
- That means you can see the characteristics of who moved and who stayed and trace the mobility patterns of who moved and those who stayed

- So Long and Ferrie have linked records for US sons and fathers, British fathers and with sons who stayed in Britain, and British fathers with sons who moved to the US
- They are going to think about migration decisions being the outcome of three equations:

$$y_{1i} = \beta_1 X_{1i} + \varepsilon_{1i} \text{ if } M_i = 1$$

$$y_{0i} = \beta_0 X_{0i} + \varepsilon_{0i} \text{ if } M_i = 0$$

$$M_i = \begin{cases} 1 & \text{if } \gamma_1 Z_i + \gamma_2 (y_{1i} - y_{0i}) + u_i \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

$$y_{1i} = \beta_1 X_{1i} + \varepsilon_{1i} \text{ if } M_i = 1$$
$$y_{0i} = \beta_0 X_{0i} + \varepsilon_{0i} \text{ if } M_i = 0$$

$$M_i = \begin{cases} 1 & \text{if } \gamma_1 Z_i + \gamma_2 (y_{1i} - y_{0i}) + u_i \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

- Here y_{ii} is the individual i's outcome in country j and X_{ii} is the individual's observable characteristics
- y_{ii} is an ordered set of occupational outcomes, from best to worst it is: white collar, farmer, skilled and semiskilled, and unskilled
- M_i equals one if the individual migrates and zero if the individual doesn't

$$y_{1i} = \beta_1 X_{1i} + \varepsilon_{1i} \text{ if } M_i = 1$$
$$y_{0i} = \beta_0 X_{0i} + \varepsilon_{0i} \text{ if } M_i = 0$$

$$M_i = \begin{cases} 1 & \text{if } \gamma_1 Z_i + \gamma_2 (y_{1i} - y_{0i}) + u_i \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

- You can estimate these equations using the observed migrants and stayers
- Once you estimate the equations, you can use the coefficients to predict what outcomes would have been if migrants stayed or stayers migrated

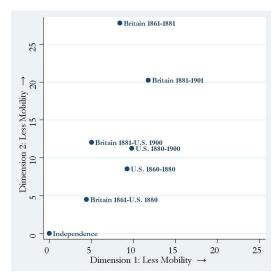
Ordered Probit Switching Regression											
1881	Mo	vers	Sta	ayers	Str	uctural P	robit (Mo	ve)			
Characteristic	β	<i>t</i> -stat.	β	<i>t</i> -stat.	β	S.E.	[90%	C.I.]			
Father's Class: 1. WO	0.56	4.39***	0.97	13.35***							
Father's Class: 2. F	0.48	2.30**	0.88	10.02***							
Father's Class: 3. SS	0.25	2.40**	0.37	6.35***							
Age	0.13	0.93	0.10	1.47	0.16	0.15	[-0.10	0.41]			
Age ²	0.00	0.85	0.00	1.44	-0.01	0.01	[-0.01	0.00]			
Father's Age	0.00	0.23	0.01	3.60***	0.00	0.01	[-0.01	0.01]			
Father in Agric.	-0.06	0.33	-0.43	6.19***							
One Servant in HH	0.34	2.40**	0.34	4.07***	0.11	0.17	[-0.17	0.40]			
2+ Servants in HH	0.42	2.48**	0.55	4.96***	0.02	0.23	[-0.35	0.40]			
Age Discrepancy	-0.02	0.30	-0.06	2.80***	0.16	0.06	[0.06	0.26]			
Eldest Child	-0.06	0.74	-0.04	0.91	-0.12	0.09	[-0.28	0.04]			
Oldest Brother in HI	H				-0.07	0.06	[-0.17	0.02]			
Children in HH					0.04	0.01	[0.02	0.06]			
Mother Employed					-0.19	0.07	[-0.31	-0.07]			
Parish ≠ Birth Parish	1				-0.04	0.04	[-0.10	0.03]			
$\hat{\mathbf{y}}_M - \hat{\mathbf{y}}_S$					-0.90	0.37	[-1.51	-0.29]			
Constant	-0.22	0.21	-0.37	0.74	-2.18	1.08	[-3.95	-0.39]			

Note: Observations: 5,025. Omitted categories are "Father's Class: 4. U," "No Servants in HH," "<2 Servants in HH," "Not Eldest Child," "Not Oldest Brother in HH," "Mother Not Employed," and "Parish=Birth Parish." Structural Probit SEs and CIs calculated by bootstrapping via data resampling with 500 repetitions.

* significant at 10%; ** significant at 5%; *** significant at 1%

Parameter	Estimate	S.E.	[90%	C.I.]
(1) $\hat{\mathbf{y}}_{M}$, Movers	1.093	0.127	[0.883	1.302]
(2) $\hat{\mathbf{y}}_{M}$, Stayers	1.028	0.138	[0.801	1.254]
(3) $\hat{\mathbf{y}}_{s}$, Movers	1.128	0.153	[0.876	1.380]
(4) $\hat{\mathbf{y}}_{s}$, Stayers	1.002	0.130	[0.787	1.216]
(5) s_M , Selection of migrants=(1)-(2)	0.065	0.027	[0.021	0.109]
(6) s _s , Selection of stayers=(4)-(3)	-0.127	0.029	[-0.174	-0.080]
(7) τ_M , Treatment Effect: Treated=(1)-(3)	-0.036	0.196	[-0.359	0.288]
(8) τ _s , Treatment Effect: Not Treated=(2)-(4)	0.026	0.186	[-0.281	0.333]
Average Treatment Effect	0.013	0.188	[-0.297	0.322]

Note: SEs and CIs are calculated by bootstrapping via data resampling with 500 repetitions.



- So it looks from Long and Ferrie like British migrants may have been positively selected
- It also looks like they experienced substantial occupational mobility when arriving in the US, even more than US sons were experiencing
- But we have a bit of a problem
- We only truly observe the impacts of migration for those who decide to migrate
- We can't really say what would have happened if a non-migrant was randomly chosen to migrate
- Time to get more ambitious with data linking and switch our focus to Norway

The Norwegian Census

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Matching Migrants

- Abramitzky et al. are going to match individuals between the 1865 Norwegian census and the 1900 Norwegian and United States censuses
- They use three different techniques:
 - Match 1: Use a standard iterative matching technique to match the population of Norwegian-born men in 1900 to their childhood households in 1865 using name, age, and country of birth
 - Match 2: Add province of birth for men who remain in Norway
 - Match 3: Restrict to men who are unique by name within a five-year age band in both censuses
- These different approaches represent tradeoffs between sample size and false matches

- By just observing migrants and non-migrants, we can learn a few things
- At most basic level, we can see occupational distributions
- Let's take a look at overall occupational distributions and then at the occupational distributions for migrants and non-migrants separately
- This will give us some (limited) insight into selection and returns to migration

Norwegians and the Roy Model

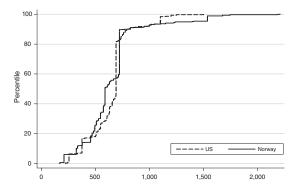


FIGURE 1. CUMULATIVE INCOME DISTRIBUTION FUNCTIONS IN THE UNITED STATES AND NORWAY IN 1900

Notes: US and Norwegian distributions contain all men aged 38 to 50 in the respective censuses of 1900. The x-axis is scaled in 1900 US dollars. Individuals are assigned the mean earnings for their occupation and are arrayed from lowest- to highest-paid occupations. The Norwegian distribution is rescaled to have the same mean as the US distribution (the actual Norwegian and US means are US\$(1900)350 and US\$(1900)643, respectively).

Rank	Occupation	Frequency	Percentage	Earnings
1	General farmers	4,189	22.26	393
2	Farmer and fisherman	1,522	8.09	321
3	Merchants and dealers	722	3.84	837
4	Fisherman	709	3.77	248
5	Husbandmen or cottars	658	3.50	114
6	Farm workers	597	3.17	175
7	Carpenters	505	2.68	312
8	Shipmasters and captains	459	2.44	298
9	Cottar and fisherman	412	2.19	321
10	Seamen	351	1.87	182
Total		10,124	53.79	

Panel B. Top ten occupations in matched sample, Norwegian-born men living in Norway in 1900

Notes: N = 18,820. Historical International Standard Classification of Occupations (HISCO) occupation categories. Annual earnings by occupation data from Statistik Centralbureau (1905) and Grytten (2007). Values reported in year 1900 dollars. Average incomes of owner-occupier farmers and fishermen are estimated using data from the Norwegian census of agriculture.

Rank	Occupation	Frequency	Percentage	Earnings	
1	Farmers and planters	1,012	35.81	691	
2	Laborers (general)	256	9.05	373	
3	Carpenters and joiners	174	6.15	630	
4	Farm laborers	101	3.57	255	
5	Painters, glaziers, and varnishers	66	2.33	624	
6	Sailors	60	2.12	467	
7	Saw and planing mill workers	42	1.49	572	
8	Machinists	39	1.38	736	
9	Railroad laborers	36	1.27	460	
10	Salesmen	32	1.13	680	
Total		1,818	64.33		

Notes: N = 2,826. Occupation data collected by hand from census manuscripts on Ancestry.com. Annual earnings by occupation data from the 1901 Cost of Living Survey reported in Preston and Haines (1991) in year 1900 dollars. Average income of owner-ocupier farmers is estimated using data from the US census of agriculture.

- Let me know if you didn't get an email confirmation for your second referee report
- Time to really focus on your data projects and thinking about research project ideas
- Today and Wednesday we are going to wrap up immigration
- Up next: education, returns to skill, and inequality
 - Gray (2013) "Taking Technology to Task: The Skill Content of Technological Change in Early 20th Century United States"
 - Parman (2011) "American Mobility and the Expansion of Public Education"

Norwegians and the Roy Model

- So in the context of Norway, it looks like we would expect negative selection of migrants
- Why? The US earnings distribution was more compressed at the time than that of Norway
- One thing to think about: could low earning Norwegians really afford the trip?
- Seems that way, the total cost of migration was about 18 percent of the annual earnings of a Norwegian farm laborer (compare this to our discussion of indentured servitude)
- How will this negative selection impact estimates of the return to migration?

• Let's start with a very naive estimate of the returns to migration:

 $ln(\text{Earnings}_i) = \alpha + \beta_1(\text{Migrant}_i) + \beta_2(\text{Age}_i) + \beta_3(\text{Age}_i)^2 + \varepsilon_i$

- $\bullet\,$ In this equation, we can think of β_1 as representing the return to migration
- Let's see what Abramitzky et al. get when running this regression

The Returns to Migration - Historical

	Dependent variable = $ln(earnings)$												
	Population (1)	Match 1 (2)	Match 2 (3)	Match 3 (4)	Weighted (5)	Match 1 Iowa data (6)	Add penalty (7)						
In US	0.609 (0.017)	0.606 (0.009)	0.644 (0.009)	0.572 (0.015)	0.641 (0.024)	0.554 (0.010)	0.466 (0.009)						
Ν	122,620	17,501	33,641	7,596	14,647	17,352	17,501						

TABLE 2-OLS REGRESSIONS OF THE RETURN TO MIGRATION FROM NORWAY TO THE UNITED STATES

Notes: Standard errors are reported in parentheses. All regressions control a quadratic in age. The first column contains a representative sample of the population of Norwegian-born men between the ages of 38–50 in 1900 from the 100 percent 1900 Norwegian census and 1 percent 1900 US census sample (IPUMS). Column 2 reports estimates from the first matched sample, which is based on an iterative matching strategy that searches first for an exact match and then for matches in a one- or two-year age band. Column 3 uses the second matched sample, which allows men to match in Norway by name, age, and province of birth. Column 4 reports estimates from the third matched sample, which instead requires that matched observations be unique within a five-year age band. Columns 5 through signed samings from the 1915 lowa census (appropriately adjusted for inflation). We lose 157 observations whose occupations do not match categories in the lowa census. In column 6, we reduce the Cost of Living earnings by 13 log points in each occupation based on the earnings penalty for Scandinavian migrants reported in Hatton and Williamson (1994). Column 7 weights the matched sample to reflect the urban status, asset holdings and occupational distribution of fathers in the full population. We lose 2,005 observations because of missing information (primarily missing data on flathers' occupations). $ln(\text{Earnings}_i) = \alpha + \beta_1(\text{Migrant}_i) + \beta_2(\text{Age}_i) + \beta_3(\text{Age}_i)^2 + \varepsilon_i$

- But there is a big problem interpreting β_1 as the return to migration
- If migrants are negatively (or positively) selected, chances are ε_i will be correlated with Migrant_i
- This is going to give us biased estimates
- So what can we do about this?

The Returns to Migration and Selection

- We can think about taking advantage of brothers in the sample
- Suppose that ε_{ij} can actually be thought of as having two components:

$$\varepsilon_{ij} = \alpha_j + \nu_{ij}$$

- Here we've added *j* to indicate individuals *i*'s household
- α_j is a household-specific component of the error term and ν_{ij} is the individual-specific component
- What happens if we run a regression with household fixed effects?

 $ln(\text{Earnings}_{ij}) = \alpha_j + \beta'_1(\text{Migrant}_{ij}) + \beta'_2(\text{Age}_{ij}) + \beta'_3(\text{Age}_{ij})^2 + \nu_{ij}$

- Including household fixed effects through α_j eliminates common household characteristics (both observed and unobserved) from the error term
- If migrants are negatively selected, we would expect β_1' to be bigger than β_1
- If migrants are positively selected, we would expect β_1' to be less than β_1
- Let's see what we get

Dependent v	nt variable = $\ln(\text{earnings})$; Coefficient on = 1 if migrant			
	Full sample, 1865	Rural, 1865	Urban, 1865	
Panel A. Unweighted				
OLS	0.545	0.607	0.384	
	(0.027)	(0.034)	(0.044)	
Within household	0.511	0.508	0.508	
	(0.035)	(0.045)	(0.057)	
Chi-squared	1.49	7.47	8.31	
p-value	0.2218	0.0063	0.0039	
N	2,655	1,823	832	
Number of migrant-stayer pairs	326	167	159	
Panel B. Weighted	0.586	0.609	0.443	
OLS	(0.029)	(0.033)	(0.067)	
Within household	0.542	0.529	0.561	
	(0.039)	(0.042)	(0.049)	
Chi-squared	2.13	4.60	5.65	
p-value	0.1441	0.0320	0.0175	
N	2,241	1,666	306	
Number of migrant-stayer pairs	269	140	129	

TABLE 3—OLS AND WITHIN-HOUSEHOLD ESTIMATES OF THE RETURN TO MIGRATION. HOUSEHOLDS WITH TWO OR MORE MEMBERS IN THE MATCHED SAMPLE

Notes: Each cell contains coefficient estimates and standard errors from regressions of ln(earnings) on a dummy variable equal to one for individuals living in the United States in 1900. Regressions also include controls for age and age squared. In each panel, the first row conducts an OLS regression for the restricted sample of households that have at least two matched members in the dataset and the second row adds household fixed effects. Panel B contains results from regressions weighted to reflect the urban status (full sample only), asset holdings, and occupational distribution of fathers in the full population. We conduct thi-squared tests of the null hypothesis that the OLS and within-household coefficients are equal.

- So have we solved our selection issue?
- Only if the problematic component was α_j, the characteristics common to both brothers
- These are certainly important (think about the role of parents) but may not capture everything that matters
- What can we do if the problem is really ν_{ij} ?
- Household fixed effects effects won't help us here



- So what can we do about ν_{ij} ?
- We can't rely on fixed effects to get rid of ν_{ij}
- What we can do is try to find some variation in the decision to migrate that is uncorrelated with ν_{ij}
- More specifically, we would like an instrument that is correlated with the decision to migrate but uncorrelated with all of the other unobserved characteristics that influence earnings
- What can we use?



Appendix Table A1—Birth Order and Number of Brothers as Instruments for Migration to the United States				
	(1)	(2)	(3)	
Panel A. First stage	Dependent variable = In US in 1900			
Number of brothers	0.016 (0.006)		0.011 (0.006)	
2nd brother		-0.000 (0.012)	_	
3rd brother		0.047 (0.019)	0.037 (0.019)	
4th or higher brother		0.076 (0.035)	0.058 (0.036)	
Panel B. OLS	Dependent	variable = ln(earnings	in 1900)	
In US in 1900	X	0.642 (0.019)	,	
Panel C. IV	Dependent variable = $\ln(\text{earnings in } 1900)$			
In US in 1900	0.669 (0.436)	0.696 (0.381)	0.668 (0.338)	
Over-ID test (p-value) N	4031	4031	0.869 4031	

Notes: Standard errors are reported in parentheses. The sample includes men in Match 1 who lived in a rural household that had some assets in 1865 and whose mother is 42 years old or younger in 1865. The regressions also include a quadratic in age and dummy variables for total number of siblings in the household (see equation (3) in the text). In column 3, we report the *p*-value from a Sargan (chi-squared) test of overidentification.

- There is one big question to think about with these results
- The returns to migration are being estimated from mean earnings by occupation
- What if migrants earn more or less within a given occupation?
- This seems like a very plausible scenario
- Let's work through this as a class in Stata

- We'll use the dataset 1940-census-income-data.dta
- These data come from the IPUMS 1% sample of the 1940 federal census
- Variables are exactly as downloaded from IPUMS, check definitions here
- As a class, we are going to figure out how to best identify whether immigrant and native earnings differ within occupational categories and much this changes our interpretation of the returns to migration
- Let's head over to Stata

- A cluster of class cancellations: March 29th, April 3rd, April 5th
- Internet may be spotty for me April 3rd through April 7th
- Consider getting a jump on data and the final research project so you can get questions answered before April 3
- Up next: education, returns to skill, and inequality
 - Gray (2013) "Taking Technology to Task: The Skill Content of Technological Change in Early 20th Century United States"
 - Parman (2011) "American Mobility and the Expansion of Public Education"

- So we have (sort of) established the returns to migration
- The next question is what inequality and mobility look like for migrants after arrival
- Do they converge with the native population in terms of earnings, education, social norms, etc.?
- Let's start by thinking about earnings convergence for the migrants themselves (their kids are a whole other story)

• Given a single cross section of data, we might think about estimating earnings convergence with the following equation:

 $ln(y_i) = \alpha + \beta_1 F B_i + \beta_2 F B_i \cdot T_i + \beta_3 F B_i \cdot T_i^2 + \gamma X_i + \varepsilon_i$

- Here ln(y_i) is log annual earnings, FB_i is an indicator for being foreign born, T_i is the number of years since migration, and X_i are observable characteristics that may influence earnings
- β₁ then tells us the earnings gap between native-born individuals and migrants when the migrant first arrives
- β₂ and β₃ then let us describe the speed at which earnings converge (or don't converge) to those of natives

TABLE 2

FOREIGN-BORN ADULT WHITE MEN, 1970					
	NATIVE	NATIVE AND FOREIGN		Foreign	
	Born	BORN		Born	
	(1)	(2)	(3)	(4)	(5)
EDUC	.07154	.07058	.07004	.07164	.05740
τ	(53.78)	(55.68)	(55.18)	(54.11)	(12.93)
	.03167	.0 3 050	.03071	.03097	.02028
<i>T</i> 2	(22.99)	(22.86)	(22.99)	(23.10)	(3.47)
	00052	00049	00050	00051	00031
LN WW	(-20.77)	(-20.45)	(-20.78)	(-20.93)	(-3.18)
	1.10335	1.10326	1.10169	1.10111	1.07151
RURALEQ1	(81.75)	(84.78)	(84.70)	(84.67)	(21.97)
	17222	16970	17080	16915	05821
SOUTHEQ1	(-20.28)	(-20.25)	(-20.39)	(-20.18)	(-1.13)
	12090	12620	12530	12389	21587
NOTMSP	(-14.17)	(-15.01)	(-14.91)	(-14.74)	(-4.38)
	30647	31078	30947	30874	34498
FOR	(-27.76)	(-28.97) .02951	(-28.86) 16359	(-28.79) .00990	(-7.66) *
(FOR) (YSM)	*	(1.75)	(-4.32) .01461	(0.18) .01555	.01500
(FOR) (YSM2)	*	*	(3.98) 00016	(4.23) 00018	(3.87) 00019
(FOR) (EDUC)	*	*	(-2.47)	(-2.79) 01619	(-2.82)
CONSTANT	- 1.03646	-1.01537	-1.00016	$(-4.23) \\ -1.02156$	78891
Observations	34,321	36,245	36,245	36,245	1,924
(N)R	.55423	.55455	.55533	.55564	.58194
R ²	.30717	.30753	.30839	.30873	.33866
Standard error	.70900	.71008	.70966		.71676

REGRESSION ANALYSIS OF EARNINGS FOR NATIVE- AND

SOURCE .--- U.S. Bureau of the Census 1972.

Note.-t-ratios in parentheses; dependent variable: natural logarithm of earnings in hundreds of dollars.

* Variable not entered.

Estimates from Chiswick (1978).

American Mobility, Spring 2019

TABLE 3

BORN WHITE MEN WITHIN COUNTRY CATEGORIES, 1970					
	Born in Mexico or Native Born of Spanish Surname*		SPEAKING DI VEL-	Foreign Born Other than English-Speaking Developed Countries [†]	
	(1)	(2)	(3)	(4)	(5)
EDUC	.03573 (4.01)	.04324 (4.28)	4 .09217 (5.70)	.05211 (10.27)	.05086 (7.06)
τ	.01211 (1.15)	.0137: (1.30)	3 .06139 (5.11)	.01147 (1.67)	.01070 (1.42)
T2	00028 (-1.62)	00030 (-1.74)	00095 (-4.49)	00018 (-1.59)	00013 (-1.33)
LN WW	1.16436 (12.47)	1.1656 (12.50)	7 1.06921 (11.39)	1.05887 (18.72)	1.05879
RURALEQ1	14442 (-1.72)	14008 (-1.67)		05025 (77)	0512 (79)
SOUTHEQ1	24159 (-3.81)	22760 (-3.56)	.12351	24956 (-4.36)	24896 (-4.34)
NOTMSP	45087 (-5.91)	45043 (-5.91)	341734	32680 (-6.16)	32709 (-6.17)
FOR	33633 (-2.55)	18680 (-1.15)		+	+
(FOR) (EDUC)	+	02403	2 ‡	:	:
(FOR) (YSM)	.02715 (2.05)	.03023	7 .01456 (1.43)	.01877 (4.15)	.01799 (3,25)
(FOR) (YSM2)	00033 (-1.38)	00038 (-1.59)		00024 (-3.09)	00024 (-3.08)
(EDUC) (YSM)	+	+	00103 (-2.06)	+	.0000
CONSTANT	73694	84163		62107	59879
Observations (N) R R ²	804 .55229 .30503	804 .55424 .30718		1,485 .56761 .32218	1,485 .56764 .32221
Standard error	.80627	.80533		.74032	.74056

REGRESSION ANALYSIS OF EARNINGS FOR ADULT FOREIGN-BORN WHITE MEN WITHIN COUNTRY CATEGORIES, 1970

SOURCE .---- U.S. Bureau of the Census 1972.

NOTE.—--ratios in parentheses: dependent variable: natural logarithm of earnings in hundreds of dollars. * For the five Southwestern states Arizona, California, Colorado, New Mexico, and Texas.

⁺ The English-speaking developed countries are Great Britain, Ireland, Canada, Australia, and New

Zealand.

* Variable not entered.

$ln(y_i) = \alpha + \beta_1 F B_i + \beta_2 F B_i \cdot T_i + \beta_3 F B_i \cdot T_i^2 + \gamma X_i + \varepsilon_i$

- These estimates often suggest rapid earnings convergence (~ 15 years) but there is a big problem with this interpretation
- If we use a single cross section, our T_i variable is picking up two very different things
 - Earnings convergence due to additional time in the US
 - Trends in cohort quality over time
- How can we get around this problem?
- One thing we can try is *repeated* cross sections

• Consider the earnings difference in the year 2010 between the cohort migrating in 1980 and the cohort migrating in 1990:

$$ln(y_{10,80}) - ln(y_{10,90})$$

- The naive interpretation is that this difference is due to the extra ten years of time in the US for the 1980 cohort
- However, consider this decomposition of the earnings difference:

$$ln(y_{10,80}) - ln(y_{10,90}) = (ln(y_{10,80}) - ln(y_{00,80})) + (ln(y_{00,80}) - ln(y_{10,90}))$$

$$ln(y_{10,80}) - ln(y_{10,90}) = (ln(y_{10,80}) - ln(y_{00,80})) + (ln(y_{00,80}) - ln(y_{10,90}))$$

- The term on the left is capturing earnings growth for the 1980 cohort from 2000 to 2010
- The term on the right is capturing difference in earnings between the 1980 and 1990 cohorts when each had been in the country for 20 years

Group and Year	Cross-Section	Within-Cohort	Across-Cohor	
of Immigration	Growth	Growth	Growth	
White				
1965–69	.0665	.0029	.0636	
	(1.46)	(.20)	(1.61)	
1960–64	.1690	0111	.1801	
	(3.33)	(17)	(4.14)	
1950–59	.0558	.0089	.0469	
	(1.80)	(.59)	(1.54)	
Black				
1965–69	.2662	.0041	.2621	
	(5.73)	(.10)	(3.35)	
1960–64	.0066	1540	.1606	
	(.10)	(-1.49)	(1.65)	
1950–59	1950–59 –.0831 (–1.33)		.1472 (1.58)	
Asian:				
1965-69	.2829	.1972	.0857	
	(20.26)	(4.62)	(2.27)	
1960-64	.1754	.1105	.0649	
	(8.48)	(1.80)	(1.30)	
1950–59	.0204	.0327	0123	
	(.96)	(.39)	(01)	

Table 3 Decomposition of Cross-Section Growth in Immigrant Earnings

Estimates from Borjas (1985).

- So do repeated cross sections solve all of our problems?
- Not quite, there is still a (potentially big) issue of return migration
- If return migration is not random, which it most certainly isn't, our remaining migrants in later census waves may be a pretty select group
- What to do? Switch to longitudinal data
- Why not do this in the first place? Same reason we've had to get creative with mobility measures
- Longitudinal data can be very hard to get your hands on and often leaves you with very small sample sizes

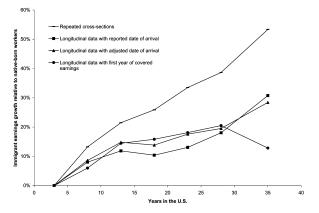
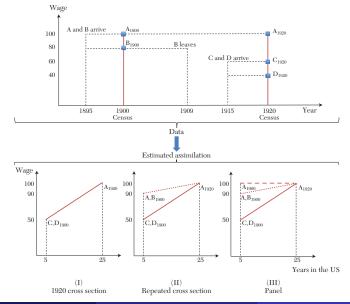


FIG. 2.—Immigrant earnings growth in repeated cross-sectional and longitudinal data. The figure plots the effect of immigrants' time in the United States on the immigrant-native earnings gap. Data are taken from estimates in table 5.



J. Parman (College of William & Mary)

American Mobility, Spring 2019

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Earnings Convergence in the Age of Mass Migration

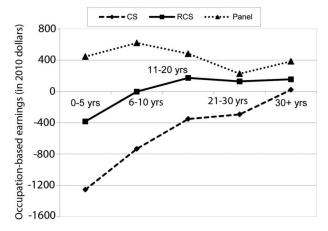


Fig. 2.

—Convergence in occupation score between immigrants and native-born workers by time spent in the United States, cross-sectional and panel data, 1900–1920. The graph plots coefficients for years spent in the United States indicators in equation (1). Note that for the panel line, we subtract the native-born dummy from the years in the United States indicators (because the omitted category in that regression is natives in the panel sample). See table 4 for coefficients and standard errors.

Earnings Convergence in the Age of Mass Migration

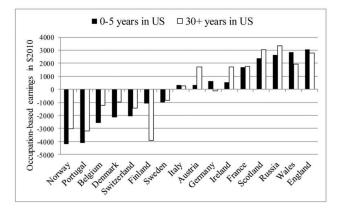
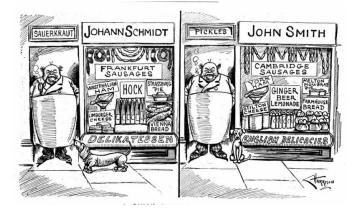


Fig. 3.

—Earnings gap between the native- and foreign-born in the panel sample: natives versus immigrants upon first arrival (0-5 years in the United States) and after time in the United States (30+ years in the United States), by country of origin. The graph reports co-efficients on the interaction between country-of-origin fixed effects and durmny variables for being in the United States for 0-5 years of rof 30+ years from regression of equation (1) in the panel sample. All coefficients for the 0-5 year interaction are significant except for Absect of Finland and Healnd.

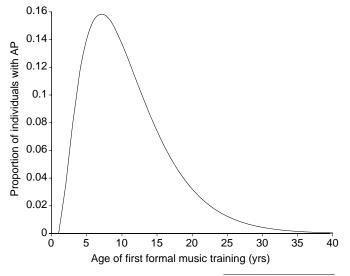
What Aids Assimilation?



- Many things might help or hinder earnings assimilation
- We'll consider a few different things:
 - Language
 - Ethnic enclaves
 - Discrimination
 - Cultural assimilation

- One critical thing may be language
- Certainly fluency in English likely helps in the US labor market
- However, we have a pretty big identification problem
- English-speaking migrants are coming from a very different set of countries than non-English-speaking migrants
- How do we tell what's due to English and what's due to other factors that differ by these countries?
- Let's take a look at the approach of Bleakley and Chin (2004)

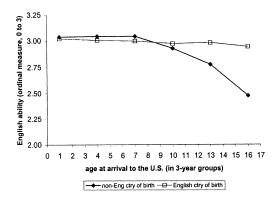


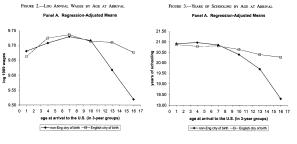


TRENDS in Cognitive Sciences

FIGURE 1.-ENGLISH-SPEAKING ABILITY BY AGE AT ARRIVAL

Panel A. Regression-Adjusted Means

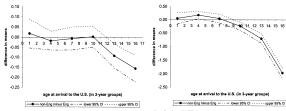




Panel B. Difference in Means

Panel B. Difference in Means

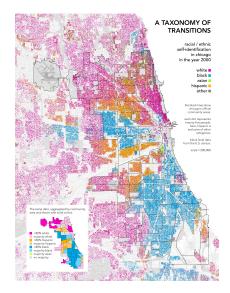
12 13 14 15 16 17



Notes: Data from 1990 IPUMS weighted by IPUMS weights. Sample size is 47,422 (composed of individuals who arrived in the United States by age 17 between 1960 and 1974 and are currently aged 25 to 38). Means have been regression-adjusted for age, race, Hispanic, and female dumnies.

Notes: Data from 1990 IPUMS weighted by IPUMS weights. Sample size is 65,214 (composed of individuals who arrived in the United States by age 17 between 1960 and 1974 and are currently aged 25 to 38). Means have been regression-adjusted for age, race, Hispanic, and female dummies.

Immigrant Enclaves and Assimilation



J. Parman (College of William & Mary)

Immigrant Enclaves and Assimilation

- A quick note on measuring segregation
- There are a variety of measures capturing different dimensions of segregation
- The dissimilarity index is a measure of evenness
- Let N_i be the number of native-born individuals in neighborhood *i* and F_i be the number of foreign-born individuals
- The dissimilarity index is then:

$$D = \frac{1}{2} \sum_{i=1}^{n} \left| \frac{F_i}{F_{total}} - \frac{N_i}{N_{total}} \right|$$

• A larger value corresponds to greater segregation

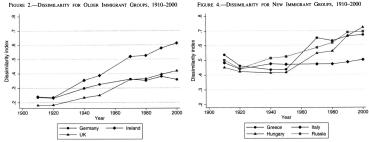


FIGURE 4.-DISSIMILARITY FOR NEW IMMIGRANT GROUPS, 1910-2000

Observations are weighted averages of statistics for immigrant communities, with weights equal to the number of immigrants in the community.

Observations are weighted averages of statistics for immigrant communities, with weights equal to the number of immigrants in the community.

Why might immigrant enclaves matter for earnings (and other types of) assimilation?

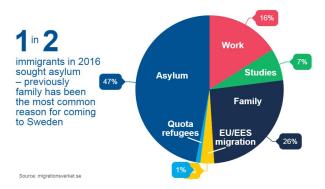
- Slower rate of acquisition of host country skills hinders moves to better jobs and resulting earnings growth
- Network effects networks present employment opportunities and a means of disseminating information about the labor market
- Spatial mismatch immigrants may be forced to segregate in an enclave far from employment opportunities
- Human capital externalities if an enclave has a high stock of human capital, this may benefit new arrivals

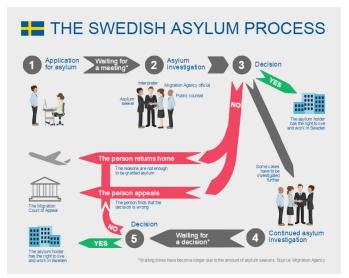
Immigrant Enclaves and Assimilation

- So immigrant enclaves might matter for all kinds of reasons
- But if a large share of immigrants live in enclaves, how do we disentangle the effects of individual characteristics versus enclave influence?
- Can we compare immigrants living in enclaves to those not living in enclaves?
- Probably not a good idea if there is self-selection into enclaves
- Let's take a look at the approach of Edin, Fredricksson and Aslund (2003)

Immigrant Enclaves and Assimilation

REASONS FOR IMMIGRATION TO SWEDEN 2016





	Initial	placement
	Enclave	No enclave
Female	.44	.45
Age	37.3	37.6
-	(7.7)	(7.4)
Years of schooling	11.3	11.7
-	(3.0)	(2.9)
Married	.63	.62
Kid ≤15 years of age	.55	.57
No. of individuals	3094	3324

TABLE I INDIVIDUAL CHARACTERISTICS BY INITIAL PLACEMENT

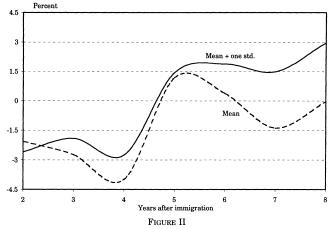
Standard deviations are in parentheses. An enclave is defined as described in the main text. Years of schooling are imputed from highest degree attained. Individuals with missing information on education were given the same number of years of schooling as those with less than nine years of schooling. All characteristics are measured eight years after immigration. The sample is restricted to those with positive earnings at that point in time.

	Full sample		(10 ye	ucation ars or ss)	High education (more than 10 years)		
	(1) OLS	(2) IV	(3) OLS	(4) IV	(5) OLS	(6) IV	
ln(ethnic group)	056	.012	053	.174	050	057	
	(.022)	(.050)	(.024)	(.088)	(.030)	(.080)	
Female	071	069	087	050	004	004	
	(.081)	(.082)	(.128)	(.132)	(.098)	(.098)	
Age	.066	.068	.079	.099	.054	.054	
	(.023)	(.022)	(.038)	(.040)	(.030)	(.031)	
Age squared (*10 ⁻²)	074	075	090	112	062	062	
• •	(.028)	(.027)	(.049)	(.052)	(.036)	(.037)	
Married	.210	.210	.289	.278	.168	.167	
	(.084)	(.084)	(.162)	(.166)	(.072)	(.073)	
Kid	027	004	115	050	.083	.081	
	(.075)	(.082)	(.132)	(.138)	(.086)	(.102)	
Married*female	049	032	226	207	.012	.011	
	(.100)	(.100)	(.153)	(.162)	(.106)	(.106)	
Kid*female	262	278	144	223	391	389	
	(.125)	(.124)	(.222)	(.229)	(.137)	(.139)	
Education missing							
and <9 years	Ref.	Ref.	Ref.	Ref.			
9–10 years	.078	.077	.097	.084			
	(.060)	(.059)	(.069)	(.070)			
High school ≤2 years	.204	.209			Ref.	Ref.	
	(.088)	(.087)					
High school >2 years	.196	.204			013	013	
0 - 0 - 0 - 0 - 0 - 0	(.070)	(.069)			(.081)	(.081)	
University <3 years	.181	.180			.006	.007	
	(.071)	(.070)			(.072)	(.072)	
University ≥3 years	.525	.526			.341	.341	
chirotony =0 years	(.081)	(.082)			(.076)	(.076)	

TABLE III BASELINE ESTIMATES—DEPENDENT VARIABLE: ln(EARNINGS)

	Full sample		(10 ye	lucation ears or ss)	High education (more than 10 years)	
	(1) IV	(2) IV	(3) IV	(4) IV	(5) IV	(6) IV
ln(ethnic group)	221 (.109)	138 (.071)	031 (.220)	.027 (.114)	315 (.161)	267 (.113)
ln(ethnic group)*ethnic inc	.044 (.015)		.039 (.040)		.047 (.019)	
ln(ethnic group)*ethnic self-employment rate		3.212 (.926)		2.964 (1.589)		4.592 (1.309)
Female	068 (.081)	071 (.080)	049 (.130)	039 (.133)	011 (.099)	018
Age	.065 (.022)	.069 (.023)	.095 (.040)	.096 (.041)	.053 (.031)	.056 (.031)
Age squared $(*10^{-2})$	072 (.027)	077 $(.028)$	108 (.052)	110 (.053)	061 (.037)	065 (.038)
Married	.204 (.088)	.203 (.084)	.287 (.168)	.290 (.162)	.147	.143
Kid	022	025 (.082)	068	080 (.134)	.057	.059
Married*female	044	(.002) 046 (.102)	(.140) 216 (.163)	236 (.152)	.002 (.108)	.008
Kid*female	(.101) 264 (.126)	(.102) 264 (.127)	(.103) 213 (.228)	(.152) 215 (.228)	(.108) 367 (.141)	(.111) 366 (.141)

TABLE IV THE "QUALITY" OF ENCLAVES—DEPENDENT VARIABLE: ln(EARNINGS)



The Reduced-Form Earnings Effect of Ethnic Concentration over Time The figure is based on Table V. The dashed line shows the effect evaluated at mean ethnic income; the solid line shows the effect evaluated at one standard deviation above the mean.

Immigrant Outcomes and Discrimination

- The immigrant enclaves raise the possibility of discrimination in the housing market influencing immigrant outcomes
- What about discrimination against immigrants in labor markets?
- If discrimination exists, it could hinder earnings convergence (or, more generally, earnings growth) after migration
- One problem is that it will be hard to find variation in discrimination within a particular ethnic group
- One solution to this problem is presented by Moser (2012)

TABLE I

Consecutive Order of Ten Ethnic Varieties in the Population of the United States according to Their Mean Rating in Ten Selected Traits

	Order of Na- tive White Americans	Order of Germans	Order of English	Order of Polish and Russian Hebrews	Order of Scandi- navians	Order of Irish	Order of French Canadians	Order of Austrian Slavs	Order of South Italians	Order of Negroes
Physical vigor Intellectual ability Self-control Moral integrity Sympathy Co-operation Leadership. Perseverance Efficiency Aspiration	3 1 3 4 6 1 1 4 1 2	2 1 1 4 2 4 1 2 4	5 4 2 2 10 3 2 3 3 3 3	8 3 5 5 5 5 6 2 4 1	1 5 4 3 9 6 5 5 5 6	46 78 1 4 38 6 5	6 8 7 8 7 7 7 7 7	7 7 8 6 7 7 8 6 8 8	9 9 9 9 9 9 9 9 9 9 9 9 9 9	10 10 10 2 10 10 10 10 10
All qualities	I	2	3	4	5	6	7	8	9	10

Survey results from Woolston, American Journal of Sociology (1916)



J. Parman (College of William & Mary)

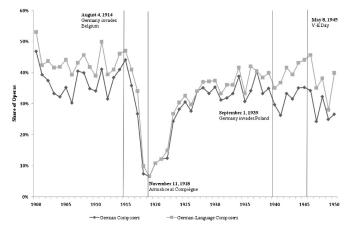


Fig. 1. The share of German-language operas from 1900 to 1950. Notes: Data on operas are collected from historical schedules of performances in the online archives of the Metropolitan Opera in New York. German-language composers include Austrian and Bohemian composers.

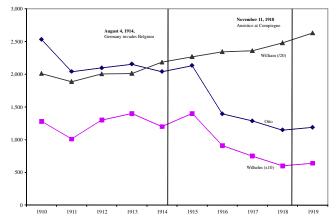


Fig. 2. Boys named Otto, Wilhelm, and William from 1910 to 1919. Notes: Data are constructed by counting the number of children with the name Otto or Wilhelm born between 1910 and 1919 and recorded in the United States Census of 1920. To scale the series in one graph, the number of Wilhelms is multiplied by 10 and the number of Williams is divided by 20.

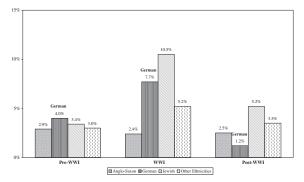


Fig. 5. Rejected applicants by ethnicity, 1883–1936. Notes: Data on names and election outcomes are collected at the archives of the NYSE. Names are matched to ethnicities by a commercial algorithm that uses linguistic rules and location-specific naming practices.

Table 3

Means of rejection rates and blackballs by ethnicity, Pre-war versus War.

	Panel A: Rejections					Panel B: Number of blackballs						
	(Non-Jewish) German		Jewish German		(Non-Jewish) German		Jewish German					
	War	Pre-War	Difference	War	Pre-War	Difference	War	Pre-War	Difference	War	Pre-War	Difference
German American	0.077 (0.016)	0.040 (0.014)	0.037 (0.021)	0.087 (0.020)	0.024 (0.018)	0.063 (0.027)	0.966 (0.179)	0.391 (0.109)	0.575 (0.210)	0.809 (0.226)	0.190 (0.137)	0.619 (0.264)
Anglo-Saxon	0.024 (0.006)	0.029 (0.005)	- 0.005 (0.008)	0.024 (0.006)	0.029 (0.005)	-0.005 (0.008)	0.304 (0.073)	0.345 (0.040)	-0.041 (0.084)	0.304 (0.071)	0.345 (0.039)	-0.041 (0.081)
Difference	0.053 (0.017)	0.011 (0.015)	0.042 [*] (0.023)	0.063 (0.021)	-0.005 (0.019)	0.068 ^{**} (0.028)	0.662 (0.194)	0.046 (0.116)	0.616 ^{***} (0.226)	0.505 (0.236)	-0.155 (0.143)	0.660 ^{**} (0.277)

Notes: The p-value for the difference in differences in rejection rates is 0.086 for non-jewish German Americans and 0.015 for jewish German Americans. The pvalue for the difference in difference in blackball is 0.007 for non-jewish German Americans and 0.017 for jewish German Americans The pto all U.S. citizens with German-sounding names. Data on admissions decisions and on the names of applicants were collected from the NYSE Archives. Names are matched to ethnicities by a commercial algorithm that takes advantage of linguistic rules and location-specific faming practices. This algorithms groups Germans with gevs together with other jewish Americans. To identify German Jews, Jewish applicants are assigned to the most frequent country of origin for immigrants with their last name in the arrival records of ships entering New York between 1850 and 1950. Standard errors in parentheses are based on a linear probability regression of rejection probabilities on ethnicities.

* p<0.10.

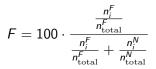
** p<0.05.

*** p<0.01.

- Moser's study points to particularly large levels of taste-based discrimination against immigrant groups
- One way to get around this taste-based discrimination is to try to appear more 'American'
- We can find evidence of this in naming practices
- Let's take a quick look at Abramitzky, Boustan and Erickson (2016)

$$R = \frac{\frac{n_i^F}{n_{\text{total}}^F}}{\frac{n_i^N}{n_{\text{total}}^N}}$$

- *R*: relative probability that a name is held by a foreign person
- n_i^F : number of foreign-born individuals with name
- n_{total}^{F} : total number of foreign-born individuals
- n_i^N : number of native-born individuals with name
- n_{total}^{N} : total number of native-born individuals



- F: foreignness index
- Goes to 100 if only foreign-born individuals have a certain name
- Goes to 0 if no foreign-born individuals have a certain name
- Goes to 50 if same percentage of foreign-born and native-born individuals have the name

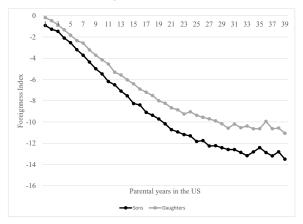
Most foreign (F-Index >0.90)	Most neutral (0.5 < F-Index < 0.52)	Most native (F-Index <0.025)	
,	A. Male names	· · · · · · · · · · · · · · · · · · ·	
Vito	Orlando	Gaylord	
Mario	Benjiman	Doyle	
Hyman	Murray	Clay	
Pasquale	Otto	Lowell	
Isidor	Theodor	Dale	
Nick	Herman	Wayne	
	B. Female names		
Sonia	Margaret	Bethany	
Antoinette	Deborah	Merlene	
Concetta	Helene	Garnet	
Johanna	Kathleen	Arlyce	
Molly	Beatrice Jo		
Carmela	Fay	Opal	

Table 1: Examples of foreign, neutral, and native names (1900-1920 birth cohorts)

Notes: Names with 100 or more observations selected for having high/lowest/most neutral F-index values in 1920 complete-count Census for the birth cohorts of 1900-20.

What Aids Assimilation?

Figure 4: Immigrants selected less foreign names for children after spending time in US, (Dependent variable = F-index)



Notes: The graph reports coefficients from estimates of Equation 1, a regression of the F-index on a set of dummy variables for years that the household head had spent in the US by the time of the child's birth. Regressions also include dummy variables for child's age and a set of household fixed effects. Data from 1920 complete-count Census. Sample includes children aged 0-18 who were born in a non-southern state and are living with their parents. Households must have a foreign-born head and the spouse (mother) must be less than 43 years old (N (sons) = 2,130,352; N (daughters) = 2,081,724).

What Aids Assimilation?

	No h	ousehold fixed e	ffects	With househo	ld fixed effects
	(1)	(2)	(3)	(4)	(5)
	Baseline	Add controls	Add F-index	Full Sample	Brothers 1-2
			at 20	-	years apart
			Panel A		
	D	ependent variat	ole: Highest gra		
F-index	-0.009***	-0.009***		-0.006***	-0.008***
	(0.0001)	(0.0001)		(0.0003)	(0.0007)
Ν	1,054,765	972,211		972,211	168,515
			Panel B		
	Deper	ndent variable:	=1 if unemploy	ed x 100 (Mear	= 9.5)
F-index	0.026***	0.026***	0.015***	0.027***	0.017***
	(0.001)	(0.002)	(0.005)	(0.003)	(0.008)
F-index at 20			0.012**		
			(0.005)		
Ν	988,383	910,936	910,936	910,936	157,531
			Panel C		
		Dep. Variable: A	nnual earnings		7)
F-index	-10.61***	-12.65***	-2.67	-6.51**	-13.67**
	(0.910)	(0.958)	(2.88)	(2.70)	(61.48)
F-index at 20			-11.04***		
			(3.00)		
Ν	673,810	620,413	620,413	620,413	107,045

Table 6: The effect of name foreignness on education, earnings and unemployment

Note: Sample includes men matched between 1920 and 1940 complete-count Censues. Men must be 3-18 in 1920, born outside the South and living at home with parents in 1920 in a household whose head was foreign-born. Panel C is further restricted to men with non-zero earnings who were not self-employed in 1940. All regressions control for a vector of dummies for child's age in 1940. Columns 2-5 control for parental years in the US and child's rank in the birth order. Columns 4-5 add household fixed effects.