Announcements

- Research projects!!!
- Officially due during our final exam period (April 29, 5pm)
- Unofficially due with enough time left for me to get grading done (May 7)
- For next week and after classes end, I'm happy to schedule individual times to meet
- On Friday, we'll use the class period as a work session for your projects, I can show you things on Stata, help with data, etc.
- Today and Wednesday we'll use a discussion of gender inequality as a way to review the course
- But first, the official launch of our class website

Launched

Economics 449: American Economic Mobility Over Two Centuries

This site displays the original work of Econ 449 students exploring the history of mobility and inequality in the Williamsburg area through a variety of sources and techniques. Each semester, students build of the work of previous cohorts, expanding the datasets and building a fuller picture of Williamsburg's economic history. Click on the images below to explore the datasets generated by the class.



The Du Bois Project

Students recreated figures from the 1900 Paris Exhibition works of W.E.B. Du Bois using historical data from Virginia and modern data from the United States. The resulting figures provide insights into the ways racial inequalities varied across the South and evolved over the twentieth century.



Neighborhood Histories

Students traced the histories of Williamsburg houses. Manuscript pages from the 1940 Federal Census reveal the neighborhood demographics from over half a century ago. Deed histories uncover the restrictive covenants that shaped the evolution of those neighborhoods.

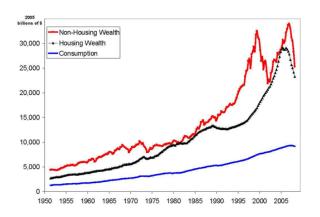


Family Histories

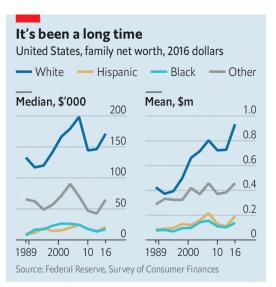
By linking Williamsburg residents from the 1920 Federal Census to the 1940 Federal census, students created an intergenerational dataset capturing occupational and geographical mobility over a time period that witnessed tremendous shocks to the national economy as well as major changes specific to the Williamsburg community.

https://jmparman.people.wm.edu/econ-449-classwebsite/econ-449-main.html

Race, Housing and Wealth



Race, Housing and Wealth



The Economist

Gender, Inequality and Mobility

Fortune 500 chief executives who are women

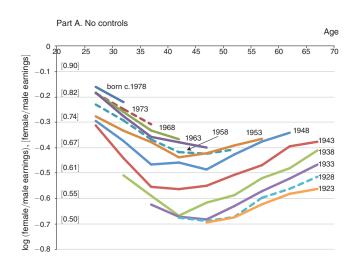


Gender, Inequality and Mobility

Fortune 500 chief executives named John



Female-Male Wage Gaps



Female-Male Wage Gaps

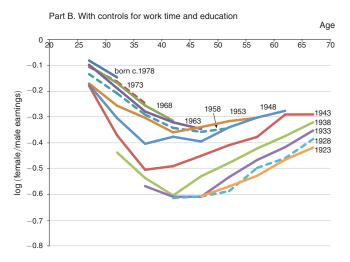


Figure 1. Relative Earnings of (Full-Time, Full-Year) College Graduate Men and Women for Synthetic Cohorts: Born 1923 to 1978

Gender, Inequality and Mobility

- For this week, we'll be thinking about gender, mobility and inequality
- Equality of opportunity and equality of outcomes varies significantly by gender
- There are a range of issues common to things we have already discussed as well as issues unique to gender
- To think about these, we'll review all of the topics we've covered in class and think about the ways in which gender enters in

What We Have Covered

- Defining and measuring inequality and mobility
- Modern versus historical estimates
- Changes in mobility and inequality over time
- Internal and international migration
- Educational and Technological Change
- Organized Labor
- Formal and informal discrimination

The Measurement of Mobility

- Let's review the measurement of mobility and rethink it in the context of female mobility
- First, our basic measures:
 - Intergenerational earnings elasticities
 - Income rank-rank correlations
 - Occupational transitions
- Do these adequately capture female mobility?
- Do they lend themselves to thinking about changes over time?

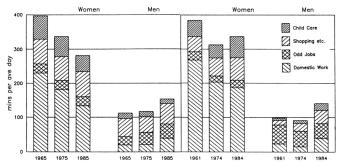


Figure 1. Estimates of the Overall Pattern of Change in Unpaid Work Time for Population Groups Aged 25-45 in the United States (left) and the United Kingdom (right)

TABLE 2-DIVORCE LAWS' IMPACT ON DIVORCE PROBABILITIES

	(1)	(2)	(3)
	Differences-in-differences	Probit	Bivariate probit
Effect of change to unilateral divorce/equitable distribution	-0.002	0.004	0.005
	(-0.27)	(0.55)	(0.43)
Effect of change to unilateral divorce/common law	-0.0003	-0.006	0.020
	(-0.03)	(-0.64)	(1.21)
Effect of change to unilateral divorce/community property	0.006	-0.004	0.001
	(0.79)	(-0.63)	(0.04)

Notes: Asymptotic t-statistics in parentheses. The dependent variable is equal to one if the woman reports herself as divorced or separated from her husband at the time of the survey. Probit estimates reported in columns (2) and (3) are $\partial P_t/\partial x_{ij} = \beta_t \cdot \Phi(x_i/\beta)$, evaluated at the mean of x_i . In addition to divorce-law variables, each regression includes an intercept, age, years of schooling, children ever born and its square, child in household under the age of six, white, year, urban residence, and state dummy variables. The sample in columns (1) and (2) consists of 73,875 married and divorced or separated women ages 18 to 55 from the 1970 and 1980 data sets. The sample in column (3) includes an additional 18,071 never-married women ages 18 to 55 used to estimate the first-stage (selection) equation of the probability of ever-heen married.

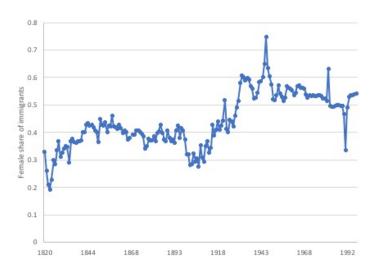
TABLE 4—DIVORCE LAWS' IMPACT ON MARRIED WOMEN'S TIME ALLOCATIONS

	Women married in either 1970 or 1980			Women married throughout sample period		
	Labor	Home	Pure	Labor	Home	Pure
	market	hours	leisure	market	hours	leisure
Effect of change to unilateral divorce/equitable distribution	-1.40	1.12	0.28	-2.44	1.64	0.80
	(-1.14)	(0.91)	(0.64)	(-1.55)	(1.00)	(0.89)
Effect of change to unilateral divorce/common law	-1.91	2.94	-1.03	-7.19	5.16	2.03
	(-0.88)	(1.36)	(-0.15)	(-2.31)	(1.61)	(0.74)
Effect of change to unilateral divorce/community property	1.34	-4.28	2.94	0.62	-6.85	6.23
	(0.74)	(-2.32)	(1.70)	(0.25)	(-2.61)	(2.33)
Sample size	2,581			1,336	_	_

Notes: Asymptotic r-statistics in parentheses. See footnote to Table 3 for other independent variables included; Census division dummy variables are included rather than state dummy variables. Tobit estimates reported as $\beta_j \cdot \Phi(\kappa/\beta)\sigma_l$, evaluated at the mean of x_i . Sample used in first three columns consists of a pooled cross section of married women in 1970 and 1980; the sample used in last three columns is restricted to women who remained married between 1970 and 1980.

The Measurement of Mobility

- Let's review the measurement of mobility and rethink it in the context of female mobility
- Now consider our data collection techniques:
 - Modern longitudinal datasets
 - Chetty-style modern data
 - Linked datasets
 - Cross-sections based on the socioeconomic content of names
- Can these techniques be applied to females?
- Are there different concerns when it comes to bias?





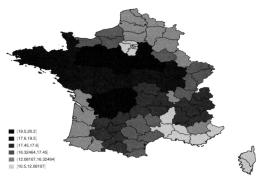


FIGURE 1. THE GEOGRAPHIC VARIATION IN MILITARY MORTALITY RATES

Notes: This map shows the geographic variation in the percentage of soldiers killed. A darker color represents a higher mortality rate (see labels for corresponding mortality rate).

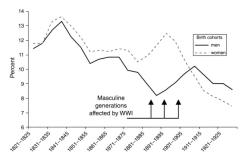


FIGURE 3. PERCENTAGE OF SINGLES AT AGE 50 BY BIRTH COHORT

Source: Roussel (1975).

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- We'll leave some time at the end of class today for course evaluations

What We Have Covered in Class

- Defining and measuring inequality and mobility
- Modern versus historical estimates
- Changes in mobility and inequality over time
- Internal and international migration
- Educational and Technological Change
- Organized Labor
- Formal and informal discrimination

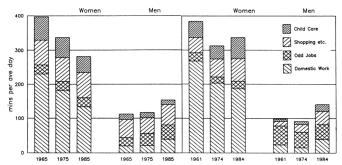


Figure 1. Estimates of the Overall Pattern of Change in Unpaid Work Time for Population Groups Aged 25–45 in the United States (left) and the United Kingdom (right)

TABLE 3—IMPACT OF MALE SCARCITY ON ASSORTATIVE MATCHING BY CLASS (OLS)
(Classes defined by own occupation)

Dependent variable	All	classes of gro	om	Excluding grooms of class 1 and 2			
	Class difference	Married down	Low- class bride	Class difference	Married down	Low- class bride	
Percent of soldiers killed	-0.020*	-0.010**	-0.017***	-0.035***	-0.017***	-0.023***	
× postwar	(0.010)	(0.004)	(0.005)	(0.013)	(0.004)	(0.006)	
Postwar	0.196 (0.146)	0.092 (0.065)	0.223*** (0.066)	0.491** (0.187)	0.240*** (0.060)	0.338*** (0.082)	
Rural	0.190**	0.032	-0.034	0.191**	0.032	-0.036	
	(0.085)	(0.029)	(0.030)	(0.091)	(0.029)	(0.033)	
Bride's age (/100)	-0.582	0.087	-0.095	-0.923*	-0.164	-0.084	
	(0.473)	(0.160)	(0.142)	(0.488)	(0.185)	(0.188)	
Groom's Age (/100)	0.355	-0.166	0.134	0.430	0.033	0.048	
	(0.377)	(0.138)	(0.127)	(0.364)	(0.130)	(0.151)	
Groom class dummies Département dummies R ² Observations	Yes	Yes	Yes	Yes	Yes	Yes	
	Yes	Yes	Yes	Yes	Yes	Yes	
	0.387	0.271	0.232	0.357	0.240	0.261	
	3,117	2,994	3,117	2,432	2,309	2,432	

Notes: The first three columns are for the full sample of grooms. The second three columns sculude grooms of classes I and 2. The dependent variable in the first and fourth columns is the class of the bride minus the class of the groom (fluss a greater class difference indicates the groom married a lower class of bride). The dependent variable in the second and fifth columns is a dummy hariable for whether the groom married a bride of lower class than himself. The dependent variable in the third and sixth columns is a dummy for the bride being of low class (5, 6, or 7). Classes are defined using own occupation. Standard errors, clustered at the département level, are presented in parentheses.

^{***}Significant at the 1 percent level.

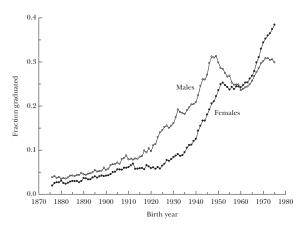
^{**}Significant at the 5 percent level.

^{*}Significant at the 10 percent level.

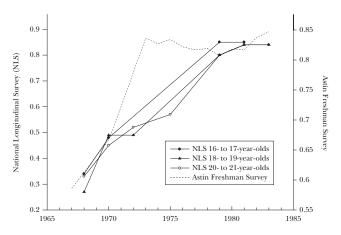
Organized Labor Well Before the Great Compression

TIME TABLE OF THE LOWELL Arranged to make the working time throughout the year average 11 hours per day. TO TAKE EFFECT SEPTEMBER 21st., 1853. The Standard time being that of the meridian of Lowell, as shown by the Regulator Clock of AMOS SANBORN, Post Office Corner, Central Street. From March 20th to September 19th, inclusive. COMMENCE WORK, at 6.30 A. M. LEAVE OFF WORK, at 6.30 P. M., except on Saturday Evenings. BREAKFAST at 6 A. M. DINNER, at 12 M. Commence Work, after dinner, 12.45 P. M. From September 20th to March 19th, inclusive. COMMENCE WORK at 7.00 A. M. LEAVE OFF WORK, at 7.00 P. M., except on Saturday Evenings. BREAKFAST at 6.30 A. M. DINNER, at 12.30 P.M. Commence Work, after dinner, 1.15 P. M. 图图11日8. From March 20th to September 19th, inclusive. Dinner Bells. From September 20th to March 19th, inclusive, Morning Bells. Evening Bells. SATURDAY EVENING BELLS. During APRIL, MAY, JUNE, JULY, and AUGUST, Ring Out, at 6.00 P. M. The remaining Saturday Evenings in the year, ring out as follows: SEPTEMBER. NOVEMBER. First Saturday, ring out 6.00 P. M. Third Saturday ring out 4.00 P. M. Third Saturday, ring out 4.25 P. M. " 5.45 " Fourth " " 3.55 " Fourth " " 4.35 " Third " ** 5.30 ** Fourth " 5.20 " DECEMBER. OCTOBER. ' First Saturday, ring out 4.45 P. M. First Saturday, ring out 3.50 P. M. .. 4.55 .. First Saturday, ring out 5.05 P. M. Second " " 3.55 " " 5.00 " " 4.55 " Third " ** 3.55 ** " 5.10 " " 4.45 " Fourth " " 4.00 " " 4.35 " Fifth " Fourth " 4 25 " First Saturday, ring out 5.25 P. M. JANUARY. Second " " 5.30 " First Saturday, ring out 4.15-P. M. First Saturday, ring out 4.10 P. M. Third " " 5.35 " Second " 4.05 " Second " 4.15 " Fourth " YARD GATES will be opened at the first stroke of the bells for entering or leaving the Mills. SPEED GATES commence hoisting three minutes before commencing work.

Figure 1 College Graduation Rates (by 35 years) for Men and Women: Cohorts Born from 1876 to 1975



 $\label{eq:Figure 6} {\it Expectations and Opinions of Female Teenagers and College Freshmen:} 1967 to 1984$



Sources: Goldin (2005), which uses the 1968 National Longitudinal Survey of Young Women (NLS68) and 1979 National Longitudinal Survey of Youth (NLSY79); Astin Oseguera, Sax, and Korn (2002).

Notes: The National Longitudinal Survey (NLS) data are the response to whether an individual stated she expected to be in the paid labor force at age 35 and are given here for white women. The Astin Freshman Survey data are the response to whether the individual disagreed with the statement "the activities of married women are best confined to the home and family." The NLS data link the averages for each age group over time. Thus, the fourteen- to fifteen-year-olds in 1978 loss sixteen-to seventeen-year-solds in 1979 and are linked to the sixteen- to seventeen-year-olds in 1979 in the NLSY79. Also, the question asked in 1968 differs somewhat from that asked in subsequent years. The Astin et al. data are for female college freshman; about 84 percent of these freshmen were 18 years old in 1967 (on December 31) and 80 percent were in 1984 (on December 31) and 80 percent were in 1984 (on December 31) and 80 percent were in 1984 (on December 31).

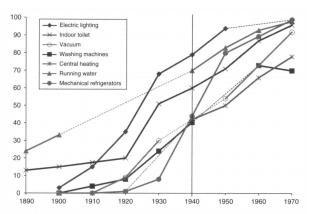
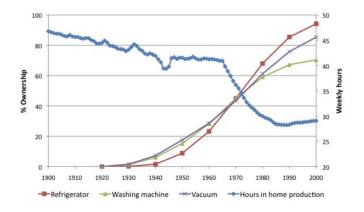
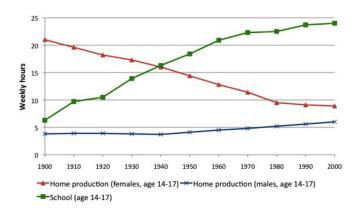


FIGURE 2. PROPORTION OF HOUSEHOLDS WITH MODERN HOUSEHOLD TECHNOLOGY, 1890-1970

Note: Dashed lines indicate linear interpolation between data points.

Source: Lebergott (1976, 260-88).

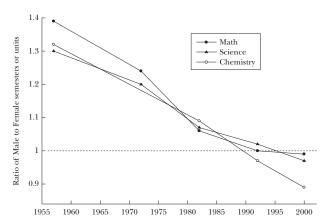




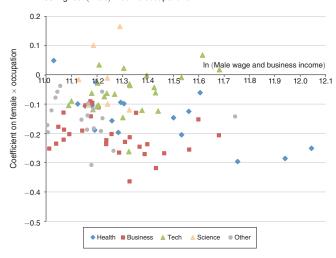




 ${\it Figure~5} \\ {\it Male-to-Female~Ratio~of~High~School~Courses~in~Math~and~Science,~1957~to~2000}$



Part C. Full-time, full-year less than 45 years old for the approximately 95 highest (male) income occupations



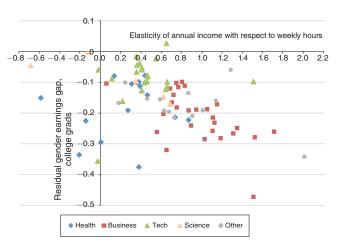


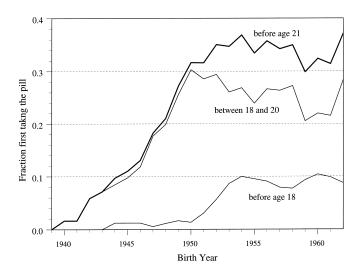
FIGURE 3. RELATIONSHIP BETWEEN THE ELASTICITY OF EARNINGS WITH RESPECT TO HOURS AND THE GENDER EARNINGS GAP



Enovid, 1960



26th Amendment, 1971



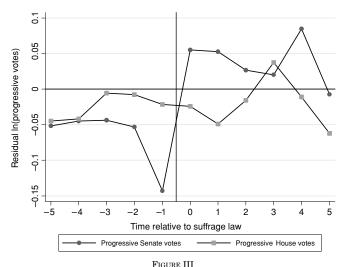
Dependent variable	$1 = First birth$ before age 22^a			1 = Before 19 ^b	1 = Before 36 ^c	Children ever born ^d	
	(1)	(2)	(3)	(4)	(5)	(6)	
Mean dependent variable		0.497		0.201	0.973	2.38	
ELA to pill	-0.071 [0.039]	-0.076 [0.039]	-0.093 [0.043]	-0.011 [0.037]	-0.001 [0.031]	-0.062 [0.086]	
Early legal access to abortion			-0.074 [0.057]	-0.086 [0.045]	-0.006 [0.006]	0.242	
ELA to pill and abortion			0.057	0.002	0.005	-0.186 [0.114]	
Fixed effects	S, C ^e	S, C, SxC ^{e,f}	S, C, SxC ^{e,f}	S, C, SxC ^{e,f}	S, C, SxC ^{e,f}	S, C, SxC ^{e,f}	
Observations Log-likelihood	$91791 \\ -62118$	91791 -61885	91791 -61866	91791 -43968	91791 -9892	91791 -145419	

TABLE IV
THE EFFECT OF EARLY ACCESS TO CONTRACEPTION ON LABOR
MARKET PARTICIPATION

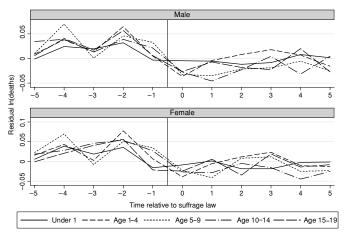
Dependent variable		1 = In the labor force						
	Mean dependent variable ^a		June					
		(1)	(2)	(3)	(4)	CPS (5)		
ELA to pill ×	0.605	0.003	0.005	-0.003	0.009	-0.048		
21-25		[0.006]	[0.006]	[0.008]	[0.009]	[0.059]		
ELA to pill ×	0.580	0.039	0.042	0.040	0.028	0.005		
26-30		[0.007]	[0.006]	[0.009]	[0.010]	[0.022]		
ELA to pill ×	0.640	0.016	0.019	0.022	0.019	0.004		
31-35		[0.006]	[0.006]	[0.009]	[0.010]	[0.021]		
ELA to pill ×	0.711	-0.002	0.002	0.004	0.007	0.001		
36-40		[0.007]	[0.006]	[0.010]	[0.008]	[0.023]		
ELA to pill \times	0.752	-0.006	-0.003	-0.007	-0.007	0.091		
41-44		[0.008]	[0.008]	[0.012]	[0.008]	[0.042]		
			R, Y, C,	R, Y, C,	R, Y, C,	S, Y, C,		
Fixed effects		R, Y, C ^b	$RxYear^c$	$RxYear^c$	$RxYear^c$	SxYear		
Age of majority								
states				X				
Abortion controls ^e					X	X		
First birth before 22 ^f						Yes		
Observations		733419	733419	245943	733419	103972		
Log likelihood		-454635	-454359	-150263	-454341	-59671		







"Progressive" Legislative Votes and Women's Suffrage Law Timing



 $\label{eq:Figure IV} F_{IGURE\ IV}$ Deaths by Age and Sex and the Timing of Suffrage Laws

Gendered Language in Teacher Reviews

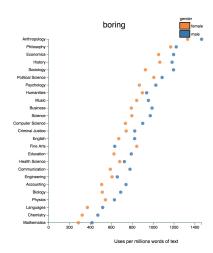
This interactive chart lets you explore the words used to describe male and female teachers in about 14 million reviews from RateMyProfessor.com.

You can enter any other word (or two-word phrase) into the box below to see how it is pit across gender and discipline: the x-axis gives how many times your term is used per million words of text (normalized against gender and field). You can also limit to just negative or positive reviews (based on the numeric ratings on the site). For some more background, see here.

Not all words have gender splits, but a surprising number do. Even things like pronouns are used quite differently by gender.

Search term(s) (case-insensitive): use commas to aggregate multiple terms





Gendered Language in Teacher Reviews